



Model Answers to Exercise 12 of December 7, 2016

<http://www.isi.ee.ethz.ch/teaching/courses/it1/>

Problem 1

Rate Distortion with Two Distortion Functions

a) We define

$$R(D_1, D_2) \triangleq \min_{P_{\hat{X}|X}(\cdot|\cdot):} I(X; \hat{X}) \quad (1)$$
$$\begin{aligned} & \mathbb{E}[d_1(X, \hat{X})] \leq D_1, \\ & \mathbb{E}[d_2(X, \hat{X})] \leq D_2 \end{aligned}$$

and we will show that $R(D_1, D_2)$ characterizes the region of achievable tuples (R, D_1, D_2) . We first prove two properties of $R(D_1, D_2)$:

Lemma 1. *The function $R(D_1, D_2)$ is nonincreasing in the pair (D_1, D_2) , i.e.,*

$$R(D'_1, D'_2) \leq R(D_1, D_2)$$

if $D'_1 \geq D_1$ and $D'_2 \geq D_2$.

Proof. Let $P_{\hat{X}|X}^*$ be a conditional PMF that minimizes the RHS of (1). For this PMF,

$$\begin{aligned} \mathbb{E}[d_1(X, \hat{X})] &\leq D_1 \leq D'_1, \\ \mathbb{E}[d_2(X, \hat{X})] &\leq D_2 \leq D'_2, \end{aligned}$$

therefore

$$R(D'_1, D'_2) = \min_{P_{\hat{X}|X}(\cdot|\cdot):} I(X; \hat{X}) \leq I(X; \hat{X}) \Big|_{P_{\hat{X}|X} = P_{\hat{X}|X}^*} = R(D_1, D_2).$$
$$\begin{aligned} & \mathbb{E}[d_1(X, \hat{X})] \leq D'_1, \\ & \mathbb{E}[d_2(X, \hat{X})] \leq D'_2 \end{aligned}$$

□

Lemma 2. *The function $R(D_1, D_2)$ is convex in the pair (D_1, D_2) , i.e.,*

$$R(\lambda D_1^{(1)} + \bar{\lambda} D_1^{(2)}, \lambda D_2^{(1)} + \bar{\lambda} D_2^{(2)}) \leq \lambda R(D_1^{(1)}, D_2^{(1)}) + \bar{\lambda} R(D_1^{(2)}, D_2^{(2)})$$

for $\lambda \in [0, 1]$ and $\bar{\lambda} = 1 - \lambda$.

Proof. Let $P_{\hat{X}|X}^{(1)}$ be a PMF that minimizes the RHS of (1) for the pair $(D_1^{(1)}, D_2^{(1)})$ and let $P_{\hat{X}|X}^{(2)}$ be a PMF that minimizes the RHS of (1) for the pair $(D_1^{(2)}, D_2^{(2)})$. Note that the mixture $\lambda P_{\hat{X}|X}^{(1)} + \bar{\lambda} P_{\hat{X}|X}^{(2)}$ is a conditional PMF satisfying

$$\begin{aligned} \mathbb{E}[d_1(X, \hat{X})] &\leq \lambda D_1^{(1)} + \bar{\lambda} D_1^{(2)}, \\ \mathbb{E}[d_2(X, \hat{X})] &\leq \lambda D_2^{(1)} + \bar{\lambda} D_2^{(2)}. \end{aligned}$$

Therefore,

$$\begin{aligned} \mathbb{R}(\lambda D_1^{(1)} + \bar{\lambda} D_1^{(2)}, \lambda D_2^{(1)} + \bar{\lambda} D_2^{(2)}) &= \min_{P_{\hat{X}|X}(\cdot|\cdot):} I(P_X, P_{\hat{X}|X}) \\ &\quad \begin{array}{l} \mathbb{E}[d_1(X, \hat{X})] \leq \lambda D_1^{(1)} + \bar{\lambda} D_1^{(2)}, \\ \mathbb{E}[d_2(X, \hat{X})] \leq \lambda D_2^{(1)} + \bar{\lambda} D_2^{(2)} \end{array} \\ &\leq I(P_X, \lambda P_{\hat{X}|X}^{(1)} + \bar{\lambda} P_{\hat{X}|X}^{(2)}) \\ &\stackrel{(i)}{\leq} \lambda I(P_X, P_{\hat{X}|X}^{(1)}) + \bar{\lambda} I(P_X, P_{\hat{X}|X}^{(2)}) \\ &= \lambda \mathbb{R}(D_1^{(1)}, D_2^{(1)}) + \bar{\lambda} \mathbb{R}(D_1^{(2)}, D_2^{(2)}), \end{aligned}$$

where (i) follows because $I(Q, W)$ is convex in W . \square

We are now ready to prove the converse, i.e., we show that $\mathbb{R} \geq \mathbb{R}(D_1, D_2)$ must hold for every encoder and decoder pair satisfying

$$\mathbb{E}[d_1(X^n, \hat{X}^n)] \leq D_1, \quad (2)$$

$$\mathbb{E}[d_2(X^n, \hat{X}^n)] \leq D_2. \quad (3)$$

This is true because

$$\begin{aligned} \mathbb{R}(D_1, D_2) &\stackrel{(i)}{\leq} \mathbb{R}(\mathbb{E}[d_1(X^n, \hat{X}^n)], \mathbb{E}[d_2(X^n, \hat{X}^n)]) \\ &\stackrel{(ii)}{=} \mathbb{R}\left(\sum_{i=1}^n \frac{1}{n} \mathbb{E}[d_1(X_i, \hat{X}_i)], \sum_{i=1}^n \frac{1}{n} \mathbb{E}[d_2(X_i, \hat{X}_i)]\right) \\ &\stackrel{(iii)}{\leq} \sum_{i=1}^n \frac{1}{n} \mathbb{R}(\mathbb{E}[d_1(X_i, \hat{X}_i)], \mathbb{E}[d_2(X_i, \hat{X}_i)]) \\ &\stackrel{(iv)}{\leq} \sum_{i=1}^n \frac{1}{n} I(X_i; \hat{X}_i) \\ &= \sum_{i=1}^n \frac{1}{n} [H(X_i) - H(X_i|\hat{X}_i)] \\ &\stackrel{(v)}{=} \frac{1}{n} H(X^n) - \frac{1}{n} \sum_{i=1}^n H(X_i|\hat{X}_i) \\ &\stackrel{(vi)}{\leq} \frac{1}{n} H(X^n) - \frac{1}{n} \sum_{i=1}^n H(X_i|\hat{X}^n, X^{i-1}) \\ &\stackrel{(vii)}{=} \frac{1}{n} H(X^n) - \frac{1}{n} H(X^n|\hat{X}^n) \\ &= \frac{1}{n} I(X^n; \hat{X}^n) \\ &\stackrel{(viii)}{\leq} \mathbb{R}, \end{aligned}$$

where (i) follows from (2), (3), and the monotonicity of $\mathsf{R}(\mathsf{D}_1, \mathsf{D}_2)$; (ii) follows because we have a single-letter distortion measure (and because expectation is linear); (iii) follows from Jensen's inequality and the convexity of $\mathsf{R}(\mathsf{D}_1, \mathsf{D}_2)$ in the pair $(\mathsf{D}_1, \mathsf{D}_2)$; (iv) follows from

$$\mathsf{R}(\mathbb{E}[d_1(X_i, \hat{X}_i)], \mathbb{E}[d_2(X_i, \hat{X}_i)]) = \min_{\substack{P_{\hat{X}|X}(\cdot|\cdot): \\ \mathbb{E}[d_1(X, \hat{X})] \leq \mathbb{E}[d_1(X_i, \hat{X}_i)], \\ \mathbb{E}[d_2(X, \hat{X})] \leq \mathbb{E}[d_2(X_i, \hat{X}_i)]}} I(P_X, P_{\hat{X}|X}) \leq I(P_{X_i}, P_{\hat{X}_i|X_i})$$

because $P_X = P_{X_i}$ and because $P_{\hat{X}_i|X_i}$ satisfies the constraints on the expected distortion; (v) follows because X_1, \dots, X_n are drawn IID; (vi) follows because conditioning does not increase entropy; (vii) follows from the chain rule; and (viii) follows from the data processing inequality and the fact that we only have 2^{nR} codewords, so $I(X^n; \hat{X}^n) \leq nR$.

We prove the direct part as follows. Fix a conditional PMF $P_{\hat{X}|X}^*$ that achieves the minimum in the RHS of (1) and compute the induced marginal distribution on \hat{X} ,

$$P_{\hat{X}}(\hat{x}) = \sum_x P_X(x) P_{\hat{X}|X}^*(\hat{x}|x), \quad \hat{x} \in \hat{\mathcal{X}}.$$

Fix $R > \mathsf{R}(\mathsf{D}_1, \mathsf{D}_2)$ and generate a codebook of size $\lceil 2^{nR} \rceil$ at random by drawing each symbol of each codeword IID according to the marginal $P_{\hat{X}}$. To describe a source sequence x^n , the encoder tries to find a codeword $\hat{x}^n(i)$ such that $(x^n, \hat{x}^n(i)) \in \mathcal{T}_{\epsilon'}^{(n)}(P_{X, \hat{X}})$ and then sends the smallest such index i . When observing i , the reconstructor produces $\hat{x}^n(i)$. Introduce a chance variable S that is one if the encoder was able to find at least one codeword $\hat{x}^n(i)$ such that $(x^n, \hat{x}^n(i)) \in \mathcal{T}_{\epsilon'}^{(n)}(P_{X, \hat{X}})$, and zero otherwise. The expected distortion for the random coding scheme can be bounded as

$$\begin{aligned} \mathbb{E}[d_1(X^n, \hat{X}^n)] &\stackrel{(i)}{=} \Pr[S = 0] \underbrace{\mathbb{E}[d_1(X^n, \hat{X}^n)|S = 0]}_{\leq d_{1, \max}} + \underbrace{\Pr[S = 1]}_{\leq 1} \underbrace{\mathbb{E}[d_1(X^n, \hat{X}^n)|S = 1]}_{\leq \mathsf{D}_1(1 + \epsilon')} \\ &\leq \Pr[S = 0] \cdot d_{1, \max} + \mathsf{D}_1(1 + \epsilon') \end{aligned} \quad (4)$$

where (i) follows from the law of total expectation; the first underbraced inequality follows because we assume bounded distortions; the second underbraced inequality is trivial; the third underbraced inequality follows because $S = 1$ implies that there exists a codeword $\hat{x}^n(i)$ such that $(x^n, \hat{x}^n(i)) \in \mathcal{T}_{\epsilon'}^{(n)}(P_{X, \hat{X}})$; and for that codeword, we have $d_1(x^n, \hat{x}^n) \leq \mathsf{D}_1(1 + \epsilon')$ as a consequence of strong typicality because $P_{\hat{X}|X}^*$ satisfies $\mathbb{E}[d_1(X, \hat{X})] \leq \mathsf{D}_1$. Similarly, we obtain

$$\mathbb{E}[d_2(X^n, \hat{X}^n)] \leq \Pr[S = 0] \cdot d_{2, \max} + \mathsf{D}_2(1 + \epsilon'). \quad (5)$$

From the lecture we know that for $R > \mathsf{R}(\mathsf{D}_1, \mathsf{D}_2) = I(P_X, P_{\hat{X}|X}^*)$ and ϵ' small enough, $\Pr[S = 0]$ tends to zero as n tends to infinity. Therefore, for an appropriate choice of ϵ' and n large enough, we obtain $\mathbb{E}[d_1(X^n, \hat{X}^n)] \leq \mathsf{D}_1 + \epsilon$ and $\mathbb{E}[d_2(X^n, \hat{X}^n)] \leq \mathsf{D}_2 + \epsilon$ from (4) and (5), respectively. Finally, we argue that we can get the same performance with a deterministic codebook. Note that (4) and (5) only depend on $\Pr[S = 0]$. By the law of total probability, we have

$$\Pr[S = 0] = \sum_{\mathcal{C}} P(\mathcal{C}) \Pr[S = 0|\mathcal{C}],$$

so there must exist a deterministic codebook \mathcal{C}^* with $\Pr[S = 0|\mathcal{C}^*] \leq \Pr[S = 0]$, and this codebook also achieves $\mathbb{E}[d_1(X^n, \hat{X}^n)] \leq \mathsf{D}_1 + \epsilon$ and $\mathbb{E}[d_2(X^n, \hat{X}^n)] \leq \mathsf{D}_2 + \epsilon$. (Since the codebook is generated independently of the source sequence, X^n is also distributed IID when conditioned on \mathcal{C}^* .)

- b) This result follows from the first part by defining $\hat{\mathcal{X}} = \hat{\mathcal{X}}_1 \times \hat{\mathcal{X}}_2$, $\hat{X} = (\hat{X}_1, \hat{X}_2)$ and setting $\tilde{d}_1(x, \hat{x}) = d_1(x, \hat{x}_1)$ and $\tilde{d}_2(x, \hat{x}) = d_2(x, \hat{x}_2)$. The two reconstructors $g_1^{(n)}$ and $g_2^{(n)}$ can be viewed as the coordinates of a single reconstructor $g^{(n)}: \{1, \dots, 2^{nR}\} \rightarrow \hat{\mathcal{X}}_1 \times \hat{\mathcal{X}}_2$. The result from the first part can then be applied to show that (R, D_1, D_2) is achievable if, and only if,

$$R \geq \min_{\substack{P_{\hat{X}|X}(\cdot|\cdot): \\ \mathbb{E}[\tilde{d}_1(X, \hat{X})] \leq D_1, \\ \mathbb{E}[\tilde{d}_2(X, \hat{X})] \leq D_2}} I(X; \hat{X}) = \min_{\substack{P_{\hat{X}_1, \hat{X}_2|X}(\cdot, \cdot|\cdot): \\ \mathbb{E}[d_1(X, \hat{X}_1)] \leq D_1, \\ \mathbb{E}[d_2(X, \hat{X}_2)] \leq D_2}} I(X; \hat{X}_1, \hat{X}_2)$$

- c) This follows by noting that if $P_{\hat{X}|X}$ achieves

$$\min_{\substack{P_{\hat{X}|X}(\cdot|\cdot): \\ \mathbb{E}[d_1(X, \hat{X})] \leq D_1, \\ \mathbb{E}[d_2(X, \hat{X})] \leq D_2}} I(X; \hat{X}),$$

then

$$P_{\hat{X}_1, \hat{X}_2|X}(\hat{x}_1, \hat{x}_2|x) = P_{\hat{X}|X}(\hat{x}_1|x) \mathbb{I}\{\hat{x}_2 = \hat{x}_1\}$$

satisfies $\mathbb{E}[d_1(X, \hat{X}_1)] \leq D_1$ and $\mathbb{E}[d_2(X, \hat{X}_2)] \leq D_2$, so that the minimum over all permissible $P_{\hat{X}_1, \hat{X}_2|X}(\cdot, \cdot|\cdot)$ can only be smaller.

Problem 2

Source–Channel Separation with Feedback

The first half of the proof proceeds along the same lines as the proof for the converse of the rate-distortion theorem; the second half of the proof is similar to the proof that feedback does not increase the capacity of a DMC. The claim is true because

$$\begin{aligned} kR(D) &\stackrel{(a)}{\leq} kR(\mathbb{E}[d(U^k, \hat{U}^k)]) \\ &\stackrel{(b)}{=} kR\left(\sum_{i=1}^k \frac{1}{k} \mathbb{E}[d(U_i, \hat{U}_i)]\right) \\ &\stackrel{(c)}{\leq} \sum_{i=1}^k R(\mathbb{E}[d(U_i, \hat{U}_i)]) \\ &\stackrel{(d)}{\leq} \sum_{i=1}^k I(U_i; \hat{U}_i) \\ &= \sum_{i=1}^k [H(U_i) - H(U_i|\hat{U}_i)] \\ &\stackrel{(e)}{=} H(U^k) - \sum_{i=1}^k H(U_i|\hat{U}_i) \\ &\stackrel{(f)}{\leq} H(U^k) - \sum_{i=1}^k H(U_i|\hat{U}^k, U^{i-1}) \\ &\stackrel{(g)}{=} H(U^k) - H(U^k|\hat{U}^k) \\ &= I(U^k; \hat{U}^k) \\ &\stackrel{(h)}{\leq} I(U^k; Y^n) \end{aligned}$$

$$\begin{aligned}
&= H(Y^n) - H(Y^n|U^k) \\
&\stackrel{(i)}{=} \sum_{i=1}^n \left[H(Y_i|Y^{i-1}) - H(Y_i|U^k, Y^{i-1}) \right] \\
&\stackrel{(j)}{\leq} \sum_{i=1}^n \left[H(Y_i) - H(Y_i|U^k, Y^{i-1}) \right] \\
&\stackrel{(k)}{=} \sum_{i=1}^n \left[H(Y_i) - H(Y_i|U^k, Y^{i-1}, X_i) \right] \\
&\stackrel{(l)}{=} \sum_{i=1}^n \left[H(Y_i) - H(Y_i|X_i) \right] \\
&= \sum_{i=1}^n I(X_i; Y_i) \\
&\stackrel{(m)}{\leq} nC,
\end{aligned}$$

where (a) follows from the monotonicity of $\mathbf{R}(\mathbf{D})$; (b) follows because we use a single-letter distortion measure; (c) follows from Jensen's inequality and the convexity of $\mathbf{R}(\mathbf{D})$; (d) follows from the definition of $\mathbf{R}(\mathbf{D})$; (e) follows because U_1, \dots, U_k are independent; (f) follows because conditioning does not increase entropy; (g) follows from the chain rule; (h) follows from the data-processing inequality as \hat{U}^k is a function of Y^n ; (i) follows from the chain rule; (j) follows because conditioning does not increase entropy; (k) follows because X_i is a function of U^k and Y^{i-1} ; (l) follows because the output Y_i of a DMC depends only on its input X_i ; and (m) follows from the definition of C .

Problem 3

Source-Channel Separation and Average Bit-Error Probability

Using the distortion function

$$d(u, \hat{u}) \triangleq I\{u \neq \hat{u}\} = \begin{cases} 0 & \text{if } u = \hat{u}, \\ 1 & \text{if } u \neq \hat{u}, \end{cases}$$

we see that

$$\frac{1}{k} \sum_{i=1}^k \Pr[U_i \neq \hat{U}_i] = \frac{1}{k} \sum_{i=1}^k \mathbb{E}[d(U_i, \hat{U}_i)] = \mathbb{E}[d(U^k, \hat{U}^k)] \geq D_{\min},$$

where D_{\min} denotes the smallest expected distortion that can be achieved. From Problem 2 we know that D_{\min} has to satisfy

$$\mathbf{R}(D_{\min}) \leq \frac{n}{k}C.$$

Since we can always achieve an expected distortion of $\frac{1}{2}$ (by setting \hat{U}_i to zero for example), we have $D_{\min} \in [0, \frac{1}{2}]$, and $\mathbf{R}(D_{\min}) = 1 - H_b(D_{\min})$ follows from the example given in the lecture (Ber(1/2) source with Hamming distortion). Combining these results, we obtain

$$1 - H_b(D_{\min}) \leq \frac{n}{k}C.$$

By the assumption that $\frac{k}{n} > C$ and by the monotonicity of the binary entropy function, we get

$$\frac{1}{k} \sum_{i=1}^k \Pr[U_i \neq \hat{U}_i] \geq D_{\min} \geq H_b^{-1}\left(1 - \frac{n}{k}C\right) > 0,$$

which shows that reliable communication above channel capacity is not possible (even with feedback, since the result from Problem 2 also holds if feedback is provided to the encoder). (Without going into the details, note that this result is stronger than the converse for channel coding given in the lecture: There it was shown that the *block error probability*, i.e., the probability that at least one bit will not be decoded correctly, is bounded away from zero. Here you have shown that the *bit error probability* is bounded away from zero, i.e., that on average a constant fraction of the bits will not be decoded correctly.)