



Strong Typicality and Covering Lemma

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Let P be a probability mass function on a finite set \mathcal{X} . For a fixed $\epsilon \in (0, 1)$, the set of strongly typical x -sequences of length n is defined as

$$\mathcal{T}_\epsilon^{(n)}(P) \triangleq \{x^n \in \mathcal{X}^n \mid \forall \xi \in \mathcal{X}: |\frac{1}{n}N(\xi|x^n) - P(\xi)| \leq \epsilon P(\xi)\}, \quad (1)$$

where $N(\xi|x^n)$ denotes the number of occurrences of the symbol ξ in the sequence x^n . Using $\delta \triangleq \epsilon H(P)$, the set $\mathcal{T}_\epsilon^{(n)}(P)$ has the following properties:

- 1) For all $x^n \in \mathcal{T}_\epsilon^{(n)}(P)$, $2^{-n(H(P)+\delta)} \leq \prod_{i=1}^n P(x_i) \leq 2^{-n(H(P)-\delta)}$.
- 2) $|\mathcal{T}_\epsilon^{(n)}(P)| \leq 2^{n(H(P)+\delta)}$.
- 3) For sufficiently large n , $|\mathcal{T}_\epsilon^{(n)}(P)| \geq (1 - \epsilon) 2^{n(H(P)-\delta)}$.
- 4) For X^n IID $\sim P$, $\lim_{n \rightarrow \infty} \Pr[X^n \in \mathcal{T}_\epsilon^{(n)}(P)] = 1$.

Let P_{XY} be a joint probability mass function on finite sets \mathcal{X} and \mathcal{Y} . For a fixed $\epsilon \in (0, 1)$, the set of jointly typical sequences of length n is defined as

$$\mathcal{T}_\epsilon^{(n)}(P_{XY}) \triangleq \{(x^n, y^n) \in \mathcal{X}^n \times \mathcal{Y}^n \mid \forall (\xi, \gamma) \in \mathcal{X} \times \mathcal{Y}: |\frac{1}{n}N((\xi, \gamma)|x^n, y^n) - P_{XY}(\xi, \gamma)| \leq \epsilon P_{XY}(\xi, \gamma)\}. \quad (2)$$

Using $\delta_{XY} \triangleq \epsilon H(P_{XY})$ and denoting the marginal probability mass functions of P_{XY} by P_X and P_Y , the set $\mathcal{T}_\epsilon^{(n)}(P_{XY})$ has the following properties:

- 5) For all $(x^n, y^n) \in \mathcal{T}_\epsilon^{(n)}(P_{XY})$, $x^n \in \mathcal{T}_\epsilon^{(n)}(P_X)$ and $y^n \in \mathcal{T}_\epsilon^{(n)}(P_Y)$.
- 6) For all $(x^n, y^n) \in \mathcal{T}_\epsilon^{(n)}(P_{XY})$, $2^{-n(H(P_{XY})+\delta_{XY})} \leq \prod_{i=1}^n P_{XY}(x_i, y_i) \leq 2^{-n(H(P_{XY})-\delta_{XY})}$.
- 7) $|\mathcal{T}_\epsilon^{(n)}(P_{XY})| \leq 2^{n(H(P_{XY})+\delta_{XY})}$.
- 8) For sufficiently large n , $|\mathcal{T}_\epsilon^{(n)}(P_{XY})| \geq (1 - \epsilon) 2^{n(H(P_{XY})-\delta_{XY})}$.
- 9) For (X^n, Y^n) IID $\sim P_{XY}$, $\lim_{n \rightarrow \infty} \Pr[(X^n, Y^n) \in \mathcal{T}_\epsilon^{(n)}(P_{XY})] = 1$.

Let $I(X; Y)$ be with respect to P_{XY} . For a fixed ϵ' satisfying $0 < \epsilon' < \epsilon$ and sufficiently large n , i.e., for $n \geq n_0$, where n_0 depends only on P_{XY} , ϵ , and ϵ' , the following holds:

- 10) For all $x^n \in \mathcal{T}_{\epsilon'}^{(n)}(P_X)$ and for Y^n IID $\sim P_Y$, $\Pr[(x^n, Y^n) \in \mathcal{T}_\epsilon^{(n)}(P_{XY})] \geq 2^{-n(I(X; Y)+4\delta_{XY})}$.

Lemma 1. For all $x^n \in \mathcal{T}_\epsilon^{(n)}(P)$, $2^{-n(H(P)+\delta)} \leq \prod_{i=1}^n P(x_i) \leq 2^{-n(H(P)-\delta)}$.

Proof. Fix $x^n \in \mathcal{T}_\epsilon^{(n)}(P)$. Observe that by (1), $N(\xi|x^n) = 0$ for all $\xi \in \mathcal{X}$ with $P(\xi) = 0$, which implies $\prod_{i=1}^n P(x_i) > 0$. Defining $\mathcal{X}' \triangleq \text{supp}(P)$, we have

$$\begin{aligned}
-\frac{1}{n} \log \prod_{i=1}^n P(x_i) &= -\frac{1}{n} \sum_{i=1}^n \log P(x_i) \\
&= -\frac{1}{n} \sum_{i=1}^n \sum_{\xi \in \mathcal{X}'} \mathbb{I}\{x_i = \xi\} \log P(\xi) \\
&= -\frac{1}{n} \sum_{\xi \in \mathcal{X}'} \sum_{i=1}^n \mathbb{I}\{x_i = \xi\} \log P(\xi) \\
&= -\frac{1}{n} \sum_{\xi \in \mathcal{X}'} N(\xi|x^n) \log P(\xi) \\
&= \sum_{\xi \in \mathcal{X}'} \frac{1}{n} N(\xi|x^n) \log \frac{1}{P(\xi)} \\
&\stackrel{(i)}{\leq} \sum_{\xi \in \mathcal{X}'} (1 + \epsilon) P(\xi) \log \frac{1}{P(\xi)} \\
&= (1 + \epsilon) H(P) \\
&= H(P) + \delta,
\end{aligned}$$

where (i) holds because $N(\xi|x^n) \leq n(1 + \epsilon)P(\xi)$ by (1) and because $\log \frac{1}{P(\xi)} \geq 0$ for all $\xi \in \mathcal{X}'$. This implies $\prod_{i=1}^n P(x_i) \geq 2^{-n(H(P)+\delta)}$. Similarly, using $N(\xi|x^n) \geq n(1 - \epsilon)P(\xi)$ in (i), we obtain

$$-\frac{1}{n} \log \prod_{i=1}^n P(x_i) \geq H(P) - \delta$$

and $\prod_{i=1}^n P(x_i) \leq 2^{-n(H(P)-\delta)}$. ■

Lemma 2. $|\mathcal{T}_\epsilon^{(n)}(P)| \leq 2^{n(H(P)+\delta)}$.

Proof. We have

$$\begin{aligned}
1 &= \sum_{x^n \in \mathcal{X}^n} \prod_{i=1}^n P(x_i) \\
&\geq \sum_{x^n \in \mathcal{T}_\epsilon^{(n)}(P)} \prod_{i=1}^n P(x_i) \\
&\stackrel{(i)}{\geq} \sum_{x^n \in \mathcal{T}_\epsilon^{(n)}(P)} 2^{-n(H(P)+\delta)} \\
&= |\mathcal{T}_\epsilon^{(n)}(P)| \cdot 2^{-n(H(P)+\delta)},
\end{aligned}$$

where (i) follows from Lemma 1. Multiplication by $2^{n(H(P)+\delta)}$ leads to the desired inequality. ■

Lemma 3. For sufficiently large n , $|\mathcal{T}_\epsilon^{(n)}(P)| \geq (1 - \epsilon) 2^{n(H(P) - \delta)}$.

Proof. Let X^n be IID $\sim P$. For sufficiently large n ,

$$\begin{aligned}
1 - \epsilon &\stackrel{(i)}{\leq} \Pr[X^n \in \mathcal{T}_\epsilon^{(n)}(P)] \\
&= \sum_{x^n \in \mathcal{T}_\epsilon^{(n)}(P)} \Pr[X^n = x^n] \\
&= \sum_{x^n \in \mathcal{T}_\epsilon^{(n)}(P)} \prod_{i=1}^n P(x_i) \\
&\stackrel{(ii)}{\leq} \sum_{x^n \in \mathcal{T}_\epsilon^{(n)}(P)} 2^{-n(H(P) - \delta)} \\
&= |\mathcal{T}_\epsilon^{(n)}(P)| \cdot 2^{-n(H(P) - \delta)},
\end{aligned}$$

where (i) follows from Lemma 4 for sufficiently large n ; and (ii) follows from Lemma 1. Multiplication by $2^{n(H(P) - \delta)}$ leads to the desired inequality. \blacksquare

Lemma 4. For X^n IID $\sim P$, $\lim_{n \rightarrow \infty} \Pr[X^n \in \mathcal{T}_\epsilon^{(n)}(P)] = 1$.

Proof. This follows from the weak law of large numbers and the union bound. For convenience, we provide a proof with explicit constants. Define

$$C \triangleq \min_{x \in \text{supp}(P)} P(x) > 0. \quad (3)$$

We show below that for all $\xi \in \mathcal{X}$,

$$\Pr \left[\left| \frac{1}{n} N(\xi | X^n) - P(\xi) \right| > \epsilon P(\xi) \right] \leq \frac{1}{n\epsilon^2 C^2}. \quad (4)$$

Consequently,

$$\begin{aligned}
\Pr[X^n \notin \mathcal{T}_\epsilon^{(n)}(P)] &= \Pr \left[\exists \xi \in \mathcal{X} : \left| \frac{1}{n} N(\xi | X^n) - P(\xi) \right| > \epsilon P(\xi) \right] \\
&\stackrel{(i)}{\leq} \sum_{\xi \in \mathcal{X}} \Pr \left[\left| \frac{1}{n} N(\xi | X^n) - P(\xi) \right| > \epsilon P(\xi) \right] \\
&\stackrel{(ii)}{\leq} \sum_{\xi \in \mathcal{X}} \frac{1}{n\epsilon^2 C^2} \\
&= \frac{|\mathcal{X}|}{n\epsilon^2 C^2}, \quad (5)
\end{aligned}$$

where (i) follows from the union bound; and (ii) follows from (4). Since the RHS of (5) tends to zero as n tends to infinity, $\Pr[X^n \notin \mathcal{T}_\epsilon^{(n)}(P)]$ must also tend to zero as n tends to infinity, which proves the claim.

It remains to show (4). For all $\xi \in \mathcal{X}$ with $P(\xi) = 0$, the LHS of (4) is zero, so (4) holds in these cases. Now, fix $\xi \in \mathcal{X}$ with $P(\xi) > 0$. Introduce indicator random variables Z_1, \dots, Z_n as follows:

$$Z_i \triangleq \begin{cases} 1 & \text{if } X_i = \xi, \\ 0 & \text{otherwise.} \end{cases}$$

Define $Z \triangleq \sum_{i=1}^n Z_i$, so $Z = N(\xi|X^n)$. By the linearity of the expectation,

$$\mu_Z \triangleq \mathbb{E}[Z] = \sum_{i=1}^n \mathbb{E}[Z_i] = \sum_{i=1}^n P(\xi) = nP(\xi).$$

Now,

$$\begin{aligned} \Pr \left[\left| \frac{1}{n} N(\xi|X^n) - P(\xi) \right| > \epsilon P(\xi) \right] &= \Pr[|Z - nP(\xi)| > n\epsilon P(\xi)] \\ &= \Pr[|Z - \mu_Z| > n\epsilon P(\xi)] \\ &\stackrel{(i)}{\leq} \frac{\text{Var}(Z)}{n^2 \epsilon^2 P(\xi)^2} \\ &\stackrel{(ii)}{\leq} \frac{1}{n\epsilon^2 P(\xi)^2} \\ &\stackrel{(iii)}{\leq} \frac{1}{n\epsilon^2 C^2}, \end{aligned}$$

where (i) follows from Chebyshev's inequality (see Exercise 1, Problem 4); (ii) holds since Z_1, \dots, Z_n are independent, so

$$\text{Var}(Z) = \text{Var} \left(\sum_{i=1}^n Z_i \right) = \sum_{i=1}^n \text{Var}(Z_i) = \sum_{i=1}^n P(\xi)(1 - P(\xi)) \leq \sum_{i=1}^n 1 = n;$$

and (iii) follows from (3). ■

Lemma 5. For all $(x^n, y^n) \in \mathcal{T}_\epsilon^{(n)}(P_{XY})$, $x^n \in \mathcal{T}_\epsilon^{(n)}(P_X)$ and $y^n \in \mathcal{T}_\epsilon^{(n)}(P_Y)$.

Proof. Fix $(x^n, y^n) \in \mathcal{T}_\epsilon^{(n)}(P_{XY})$. For all $\xi \in \mathcal{X}$,

$$\begin{aligned} N(\xi|x^n) &= \sum_{\gamma \in \mathcal{Y}} N((\xi, \gamma)|x^n, y^n) \\ &\stackrel{(i)}{\leq} \sum_{\gamma \in \mathcal{Y}} n(1 + \epsilon) P_{XY}(\xi, \gamma) \\ &= n(1 + \epsilon) P_X(\xi), \end{aligned}$$

where (i) follows from (2). Similarly, $N(\xi|x^n) \geq n(1 - \epsilon) P_X(\xi)$ for all $\xi \in \mathcal{X}$. Together, this implies $|\frac{1}{n} N(\xi|x^n) - P_X(\xi)| \leq \epsilon P_X(\xi)$ for all $\xi \in \mathcal{X}$, so $x^n \in \mathcal{T}_\epsilon^{(n)}(P_X)$. Analogously, $y^n \in \mathcal{T}_\epsilon^{(n)}(P_Y)$. ■

Lemma 6. For all $(x^n, y^n) \in \mathcal{T}_\epsilon^{(n)}(P_{XY})$, $2^{-n(H(P_{XY}) + \delta_{XY})} \leq \prod_{i=1}^n P_{XY}(x_i, y_i) \leq 2^{-n(H(P_{XY}) - \delta_{XY})}$.

Lemma 7. $|\mathcal{T}_\epsilon^{(n)}(P_{XY})| \leq 2^{n(H(P_{XY}) + \delta_{XY})}$.

Lemma 8. For sufficiently large n , $|\mathcal{T}_\epsilon^{(n)}(P_{XY})| \geq (1 - \epsilon) 2^{n(H(P_{XY}) - \delta_{XY})}$.

Lemma 9. For $(X^n, Y^n) \text{ IID } \sim P_{XY}$, $\lim_{n \rightarrow \infty} \Pr[(X^n, Y^n) \in \mathcal{T}_\epsilon^{(n)}(P_{XY})] = 1$.

Proof of Lemmas 6–9. They follow from applying Lemmas 1–4 to the set $\mathcal{X}' \triangleq (\mathcal{X} \times \mathcal{Y})$. ■

Lemma 10. For a fixed ϵ' satisfying $0 < \epsilon' < \epsilon$ and sufficiently large n , the following holds: For all $x^n \in \mathcal{T}_{\epsilon'}^{(n)}(P_X)$ and for Y^n IID $\sim P_Y$, $\Pr[(x^n, Y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})] \geq 2^{-n(I(X;Y)+4\delta_{XY})}$.

Proof. We have

$$\begin{aligned}
\Pr[(x^n, Y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})] &= \sum_{y^n : (x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})} \Pr[Y^n = y^n] \\
&= \sum_{y^n : (x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})} \prod_{i=1}^n P_Y(y_i) \\
&\stackrel{(i)}{\geq} \sum_{y^n : (x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})} 2^{-n(H(P_Y) + \epsilon H(P_Y))} \\
&\stackrel{(ii)}{\geq} \sum_{y^n : (x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})} 2^{-n(H(P_Y) + \delta_{XY})} \\
&= |\{y^n \in \mathcal{Y}^n : (x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})\}| \cdot 2^{-n(H(P_Y) + \delta_{XY})} \\
&\stackrel{(iii)}{\geq} 2^{n(H(P_{XY}) - H(P_X) - 3\delta_{XY})} \cdot 2^{-n(H(P_Y) + \delta_{XY})} \\
&= 2^{-n(H(P_X) + H(P_Y) - H(P_{XY}) + 4\delta_{XY})} \\
&= 2^{-n(I(X;Y) + 4\delta_{XY})},
\end{aligned}$$

where (i) follows from Lemma 1 because $y^n \in \mathcal{T}_{\epsilon}^{(n)}(P_Y)$ by Lemma 5; (ii) holds because $H(P_Y) \leq H(P_{XY})$; and (iii) follows from Lemma 12 below. \blacksquare

Lemma 11. For a fixed ϵ' satisfying $0 < \epsilon' < \epsilon$ and sufficiently large n , the following holds: For all $x^n \in \mathcal{T}_{\epsilon'}^{(n)}(P_X)$ and for $Y^n \sim \prod_{i=1}^n P_{Y|X}(y_i|x_i)$, $\lim_{n \rightarrow \infty} \Pr[(x^n, Y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})] = 1$.

Proof. Define

$$C \triangleq \min_{(x,y) \in \text{supp}(P_{XY})} P_{XY}(x, y) > 0. \quad (6)$$

We show below that for all $x^n \in \mathcal{T}_{\epsilon'}^{(n)}(P_X)$ and for all $(\xi, \gamma) \in \mathcal{X} \times \mathcal{Y}$,

$$\Pr\left[\left|\frac{1}{n}N((\xi, \gamma)|x^n, Y^n) - P_{XY}(\xi, \gamma)\right| > \epsilon P_{XY}(\xi, \gamma)\right] \leq \frac{1}{n(\epsilon - \epsilon')^2 C^2}. \quad (7)$$

Consequently, for all $x^n \in \mathcal{T}_{\epsilon'}^{(n)}(P_X)$,

$$\begin{aligned}
\Pr[(x^n, Y^n) \notin \mathcal{T}_{\epsilon}^{(n)}(P_{XY})] &= \Pr\left[\exists(\xi, \gamma) \in \mathcal{X} \times \mathcal{Y} : \left|\frac{1}{n}N((\xi, \gamma)|x^n, Y^n) - P_{XY}(\xi, \gamma)\right| > \epsilon P_{XY}(\xi, \gamma)\right] \\
&\stackrel{(i)}{\leq} \sum_{(\xi, \gamma) \in \mathcal{X} \times \mathcal{Y}} \Pr\left[\left|\frac{1}{n}N((\xi, \gamma)|x^n, Y^n) - P_{XY}(\xi, \gamma)\right| > \epsilon P_{XY}(\xi, \gamma)\right] \\
&\stackrel{(ii)}{\leq} \sum_{(\xi, \gamma) \in \mathcal{X} \times \mathcal{Y}} \frac{1}{n(\epsilon - \epsilon')^2 C^2} \\
&= \frac{|\mathcal{X}| \cdot |\mathcal{Y}|}{n(\epsilon - \epsilon')^2 C^2}, \quad (8)
\end{aligned}$$

where (i) follows from the union bound; and (ii) follows from (7). Since the RHS of (8) tends to zero as n tends to infinity, $\Pr[(x^n, Y^n) \notin \mathcal{T}_{\epsilon}^{(n)}(P_{XY})]$ must also tend to zero as n tends to infinity, which proves the claim.

It remains to show (7). Fix $x^n \in \mathcal{T}_{\epsilon'}^{(n)}(P_X)$, $\xi \in \mathcal{X}$, and $\gamma \in \mathcal{Y}$. We first treat the case $P_{XY}(\xi, \gamma) = 0$.

- If $P_X(\xi) = 0$, then $N(\xi|x^n) = 0$, so $\Pr[N((\xi, \gamma)|x^n, Y^n) = 0] = 1$;
- otherwise, $P_X(\xi) > 0$ and $P_{Y|X}(\gamma|\xi) = 0$, so $\Pr[N((\xi, \gamma)|x^n, Y^n) = 0] = 1$, too.

Since $\Pr[N((\xi, \gamma)|x^n, Y^n) = 0] = 1$, the LHS of (7) is zero, so (7) holds in this case. Assume from now on that $P_{XY}(\xi, \gamma) > 0$, which implies $P_X(\xi) > 0$. Define the set $\mathcal{A} \triangleq \{i \in \{1, \dots, n\} \mid x_i = \xi\}$. Define $k \triangleq |\mathcal{A}|$. Observe that $k = N(\xi|x^n)$. By Lemma 5, $x^n \in \mathcal{T}_{\epsilon'}^{(n)}(P_X)$, so

$$n(1 - \epsilon')P_X(\xi) \leq k \leq n(1 + \epsilon')P_X(\xi). \quad (9)$$

In particular, $k > 0$ because $P_X(\xi) > 0$ and $0 < \epsilon' < \epsilon < 1$. Label the elements of \mathcal{A} by $\alpha_1, \dots, \alpha_k$. Introduce indicator random variables Z_1, \dots, Z_k as follows:

$$Z_i \triangleq \begin{cases} 1 & \text{if } Y_{\alpha_i} = \gamma, \\ 0 & \text{otherwise.} \end{cases}$$

Define $Z \triangleq \sum_{i=1}^k Z_i$, so $Z = N((\xi, \gamma)|x^n, Y^n)$. By the linearity of the expectation,

$$\mu_Z \triangleq \mathbb{E}[Z] = \sum_{i=1}^k \mathbb{E}[Z_i] = \sum_{i=1}^k P_{Y|X}(\gamma|\xi) = kP_{Y|X}(\gamma|\xi).$$

Observe that

$$\begin{aligned} \mu_Z - nP_{XY}(\xi, \gamma) &= kP_{Y|X}(\gamma|\xi) - nP_{XY}(\xi, \gamma) \\ &\stackrel{(i)}{\leq} n(1 + \epsilon')P_X(\xi)P_{Y|X}(\gamma|\xi) - nP_{XY}(\xi, \gamma) \\ &= n(1 + \epsilon')P_{XY}(\xi, \gamma) - nP_{XY}(\xi, \gamma) \\ &= n\epsilon'P_{XY}(\xi, \gamma), \end{aligned}$$

where (i) follows from (9). Similarly, we obtain $\mu_Z - nP_{XY}(\xi, \gamma) \geq -n\epsilon'P_{XY}(\xi, \gamma)$. Together, these inequalities imply

$$|\mu_Z - nP_{XY}(\xi, \gamma)| \leq n\epsilon'P_{XY}(\xi, \gamma). \quad (10)$$

Observe that if $|Z - \mu_Z| \leq n(\epsilon - \epsilon')P_{XY}(\xi, \gamma)$, then

$$\begin{aligned} |Z - nP_{XY}(\xi, \gamma)| &= |Z - \mu_Z + \mu_Z - nP_{XY}(\xi, \gamma)| \\ &\stackrel{(i)}{\leq} |Z - \mu_Z| + |\mu_Z - nP_{XY}(\xi, \gamma)| \\ &\stackrel{(ii)}{\leq} n(\epsilon - \epsilon')P_{XY}(\xi, \gamma) + n\epsilon'P_{XY}(\xi, \gamma) \\ &= n\epsilon P_{XY}(\xi, \gamma), \end{aligned}$$

where (i) follows from the triangle inequality; and (ii) follows from the above condition and from (10). Now,

$$\begin{aligned} \Pr \left[\left| \frac{1}{n} N((\xi, \gamma)|x^n, Y^n) - P_{XY}(\xi, \gamma) \right| > \epsilon P_{XY}(\xi, \gamma) \right] &= \Pr[|Z - nP_{XY}(\xi, \gamma)| > n\epsilon P_{XY}(\xi, \gamma)] \\ &\stackrel{(i)}{\leq} \Pr[|Z - \mu_Z| > n(\epsilon - \epsilon')P_{XY}(\xi, \gamma)] \\ &\stackrel{(ii)}{\leq} \frac{\text{Var}(Z)}{n^2(\epsilon - \epsilon')^2 P_{XY}(\xi, \gamma)^2} \end{aligned}$$

$$\begin{aligned}
& \stackrel{\text{(iii)}}{\leq} \frac{1}{n(\epsilon - \epsilon')^2 P_{XY}(\xi, \gamma)^2} \\
& \stackrel{\text{(iv)}}{\leq} \frac{1}{n(\epsilon - \epsilon')^2 C^2},
\end{aligned}$$

where (i) follows from the above observation; (ii) follows from Chebyshev's inequality; (iii) holds because Z_1, \dots, Z_k are independent, so

$$\text{Var}(Z) = \text{Var}\left(\sum_{i=1}^k Z_i\right) = \sum_{i=1}^k \text{Var}(Z_i) = \sum_{i=1}^k P_{Y|X}(\gamma|\xi)(1 - P_{Y|X}(\gamma|\xi)) \leq \sum_{i=1}^k 1 = k \leq n;$$

and (iv) follows from (6). ■

Lemma 12. *For a fixed ϵ' satisfying $0 < \epsilon' < \epsilon$ and sufficiently large n , the following holds: For all $x^n \in \mathcal{T}_{\epsilon'}^{(n)}(P_X)$, $|\{y^n \in \mathcal{Y}^n : (x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})\}| \geq 2^{n(H(P_{XY}) - H(P_X) - 3\delta_{XY})}$.*

Proof. Note that $\delta_{XY} = 0$ is only possible if $H(P_{XY}) = 0$; in that case, the claim of the lemma can be easily verified. Assume $\delta_{XY} > 0$ from now on. Let n be sufficiently large. Fix $x^n \in \mathcal{T}_{\epsilon'}^{(n)}(P_X)$. Let $Y^n \sim \prod_{i=1}^n P_{Y|X}(y_i|x_i)$. Then,

$$\begin{aligned}
2^{-n\delta_{XY}} & \stackrel{\text{(i)}}{\leq} \frac{1}{2} \\
& \stackrel{\text{(ii)}}{\leq} \Pr[(x^n, Y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})] \\
& = \sum_{y^n : (x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})} \Pr[Y^n = y^n] \\
& = \sum_{y^n : (x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})} \prod_{i=1}^n P_{Y|X}(y_i|x_i) \\
& = \sum_{y^n : (x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})} \prod_{i=1}^n \frac{P_{XY}(x_i, y_i)}{P_X(x_i)} \\
& = \sum_{y^n : (x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})} \frac{\prod_{i=1}^n P_{XY}(x_i, y_i)}{\prod_{i=1}^n P_X(x_i)} \\
& \stackrel{\text{(iii)}}{\leq} \sum_{y^n : (x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})} \frac{2^{-n(H(P_{XY}) - \delta_{XY})}}{2^{-n(H(P_X) + \epsilon H(P_X))}} \\
& \stackrel{\text{(iv)}}{\leq} \sum_{y^n : (x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})} \frac{2^{-n(H(P_{XY}) - \delta_{XY})}}{2^{-n(H(P_X) + \delta_{XY})}} \\
& = |\{y^n \in \mathcal{Y}^n : (x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})\}| \cdot 2^{-n(H(P_{XY}) - H(P_X) - 2\delta_{XY})},
\end{aligned}$$

where (i) holds for sufficiently large n because $\delta_{XY} > 0$; (ii) follows from Lemma 11 for sufficiently large n ; (iii) follows from Lemma 6 because $(x^n, y^n) \in \mathcal{T}_{\epsilon}^{(n)}(P_{XY})$ and from Lemma 1 because $x^n \in \mathcal{T}_{\epsilon'}^{(n)}(P_X)$ by Lemma 5; and (iv) holds because $H(P_X) \leq H(P_{XY})$. Multiplication by $2^{n(H(P_{XY}) - H(P_X) - 2\delta_{XY})}$ leads to the desired inequality. ■