



## Model Answers to Exercise 12 of December 6, 2017

<http://www.isi.ee.ethz.ch/teaching/courses/it1.html>

### Problem 1

### *Rate Distortion with Two Distortion Functions*

a) We first prove two properties of

$$R^I(D_1, D_2) \triangleq \min_{\substack{P_{\hat{X}|X}: \\ E[d_1(X, \hat{X})] \leq D_1, \\ E[d_2(X, \hat{X})] \leq D_2}} I(X; \hat{X}) = \min_{\substack{P_{\hat{X}|X}: \\ E[d_1(X, \hat{X})] \leq D_1, \\ E[d_2(X, \hat{X})] \leq D_2}} I(P_X, P_{\hat{X}|X}).$$

**Lemma 1.** *The function  $R^I(D_1, D_2)$  is nonincreasing in the pair  $(D_1, D_2)$ , i.e.,*

$$R^I(D'_1, D'_2) \leq R^I(D_1, D_2)$$

*if  $D'_1 \geq D_1$  and  $D'_2 \geq D_2$ .*

*Proof.* Fix  $D_1, D_2, D'_1 \geq D_1$ , and  $D'_2 \geq D_2$ . Let  $P_{\hat{X}|X}^*$  be a conditional PMF that achieves the minimum in the definition of  $R^I(D_1, D_2)$ . For this PMF,

$$\begin{aligned} R^I(D_1, D_2) &= I(P_X, P_{\hat{X}|X}^*), \\ E[d_1(X, \hat{X})] &\leq D_1 \leq D'_1, \\ E[d_2(X, \hat{X})] &\leq D_2 \leq D'_2. \end{aligned}$$

Consequently,

$$R^I(D'_1, D'_2) = \min_{\substack{P_{\hat{X}|X}: \\ E[d_1(X, \hat{X})] \leq D'_1, \\ E[d_2(X, \hat{X})] \leq D'_2}} I(P_X, P_{\hat{X}|X}) \stackrel{(i)}{\leq} I(P_X, P_{\hat{X}|X}^*) = R^I(D_1, D_2),$$

where (i) holds because  $P_{\hat{X}|X}^*$  satisfies  $E[d_1(X, \hat{X})] \leq D'_1$  and  $E[d_2(X, \hat{X})] \leq D'_2$ . ■

**Lemma 2.** *The function  $R^I(D_1, D_2)$  is convex in the pair  $(D_1, D_2)$ , i.e.,*

$$R^I(\lambda D_1^{(1)} + \bar{\lambda} D_1^{(2)}, \lambda D_2^{(1)} + \bar{\lambda} D_2^{(2)}) \leq \lambda R^I(D_1^{(1)}, D_2^{(1)}) + \bar{\lambda} R^I(D_1^{(2)}, D_2^{(2)})$$

*for  $\lambda \in [0, 1]$  and  $\bar{\lambda} = 1 - \lambda$ .*

*Proof.* Let  $P_{\hat{X}|X}^{(1)}$  be a conditional PMF that achieves the minimum in the definition of  $\mathsf{R}^I(\mathsf{D}_1^{(1)}, \mathsf{D}_2^{(1)})$ , and let  $P_{\hat{X}|X}^{(2)}$  be a conditional PMF that achieves the minimum in the definition of  $\mathsf{R}^I(\mathsf{D}_1^{(2)}, \mathsf{D}_2^{(2)})$ . Note that the mixture  $\lambda P_{\hat{X}|X}^{(1)} + \bar{\lambda} P_{\hat{X}|X}^{(2)}$  is a conditional PMF satisfying

$$\begin{aligned}\mathbb{E}[d_1(X, \hat{X})] &\leq \lambda \mathsf{D}_1^{(1)} + \bar{\lambda} \mathsf{D}_1^{(2)}, \\ \mathbb{E}[d_2(X, \hat{X})] &\leq \lambda \mathsf{D}_2^{(1)} + \bar{\lambda} \mathsf{D}_2^{(2)}.\end{aligned}$$

Consequently,

$$\begin{aligned}\mathsf{R}^I(\lambda \mathsf{D}_1^{(1)} + \bar{\lambda} \mathsf{D}_1^{(2)}, \lambda \mathsf{D}_2^{(1)} + \bar{\lambda} \mathsf{D}_2^{(2)}) &= \min_{P_{\hat{X}|X}:} I(P_X, P_{\hat{X}|X}) \\ &\quad \begin{array}{l} \mathbb{E}[d_1(X, \hat{X})] \leq \lambda \mathsf{D}_1^{(1)} + \bar{\lambda} \mathsf{D}_1^{(2)}, \\ \mathbb{E}[d_2(X, \hat{X})] \leq \lambda \mathsf{D}_2^{(1)} + \bar{\lambda} \mathsf{D}_2^{(2)} \end{array} \\ &\leq I(P_X, \lambda P_{\hat{X}|X}^{(1)} + \bar{\lambda} P_{\hat{X}|X}^{(2)}) \\ &\stackrel{(i)}{\leq} \lambda I(P_X, P_{\hat{X}|X}^{(1)}) + \bar{\lambda} I(P_X, P_{\hat{X}|X}^{(2)}) \\ &= \lambda \mathsf{R}^I(\mathsf{D}_1^{(1)}, \mathsf{D}_2^{(1)}) + \bar{\lambda} \mathsf{R}^I(\mathsf{D}_1^{(2)}, \mathsf{D}_2^{(2)}),\end{aligned}$$

where (i) holds because  $I(Q, W)$  is convex in  $W$ . ■

We are now ready to prove the converse, i.e., we show that  $\mathsf{R} \geq \mathsf{R}^I(\mathsf{D}_1, \mathsf{D}_2)$  must hold for every encoder and decoder pair satisfying

$$\mathbb{E}[d_1(X^n, \hat{X}^n)] \leq \mathsf{D}_1, \tag{1}$$

$$\mathbb{E}[d_2(X^n, \hat{X}^n)] \leq \mathsf{D}_2. \tag{2}$$

(We cheat a bit and neglect  $\delta$  here, i.e., we omit a continuity argument.) Indeed,

$$\begin{aligned}\mathsf{R}^I(\mathsf{D}_1, \mathsf{D}_2) &\stackrel{(i)}{\leq} \mathsf{R}^I(\mathbb{E}[d_1(X^n, \hat{X}^n)], \mathbb{E}[d_2(X^n, \hat{X}^n)]) \\ &\stackrel{(ii)}{=} \mathsf{R}^I\left(\sum_{i=1}^n \frac{1}{n} \mathbb{E}[d_1(X_i, \hat{X}_i)], \sum_{i=1}^n \frac{1}{n} \mathbb{E}[d_2(X_i, \hat{X}_i)]\right) \\ &\stackrel{(iii)}{\leq} \sum_{i=1}^n \frac{1}{n} \mathsf{R}^I(\mathbb{E}[d_1(X_i, \hat{X}_i)], \mathbb{E}[d_2(X_i, \hat{X}_i)]) \\ &\stackrel{(iv)}{\leq} \sum_{i=1}^n \frac{1}{n} I(X_i; \hat{X}_i) \\ &= \sum_{i=1}^n \frac{1}{n} [H(X_i) - H(X_i|\hat{X}_i)] \\ &\stackrel{(v)}{=} \frac{1}{n} H(X^n) - \frac{1}{n} \sum_{i=1}^n H(X_i|\hat{X}_i) \\ &\stackrel{(vi)}{\leq} \frac{1}{n} H(X^n) - \frac{1}{n} \sum_{i=1}^n H(X_i|\hat{X}^n, X^{i-1}) \\ &\stackrel{(vii)}{=} \frac{1}{n} H(X^n) - \frac{1}{n} H(X^n|\hat{X}^n) \\ &= \frac{1}{n} I(X^n; \hat{X}^n) \\ &\stackrel{(viii)}{\leq} \mathsf{R},\end{aligned}$$

where (i) follows from (1), (2), and the monotonicity of  $R^I(D_1, D_2)$ ; (ii) holds because we have a single-letter distortion measure (and because expectation is linear); (iii) follows from Jensen's inequality because  $R^I(D_1, D_2)$  is convex in the pair  $(D_1, D_2)$ ; (iv) follows from

$$R^I(\mathbb{E}[d_1(X_i, \hat{X}_i)], \mathbb{E}[d_2(X_i, \hat{X}_i)]) = \min_{\substack{P_{\hat{X}_i|X_i}: \\ \mathbb{E}[d_1(X, \hat{X})] \leq \mathbb{E}[d_1(X_i, \hat{X}_i)], \\ \mathbb{E}[d_2(X, \hat{X})] \leq \mathbb{E}[d_2(X_i, \hat{X}_i)]}} I(P_X, P_{\hat{X}_i|X_i}) \leq I(P_{X_i}, P_{\hat{X}_i|X_i})$$

because  $P_X = P_{X_i}$  and because  $P_{\hat{X}_i|X_i}$  satisfies the constraints on the expected distortion; (v) holds since  $X_1, \dots, X_n$  are IID; (vi) holds because conditioning does not increase entropy; (vii) follows from the chain rule; and (viii) follows from the data processing inequality and the fact that we have only  $2^{nR}$  codewords, so  $I(X^n; \hat{X}^n) \leq nR$ .

We prove the direct part as follows. Fix  $D_1$  and  $D_2$ . Let  $P_{\hat{X}_i|X_i}^*$  be a conditional PMF that achieves the minimum in the definition of  $R^I(D_1, D_2)$ . Then, for all  $\hat{x} \in \hat{\mathcal{X}}$ ,

$$P_{\hat{X}}(\hat{x}) = \sum_x P_X(x) P_{\hat{X}_i|X_i}^*(\hat{x}|x).$$

Fix  $R > R^I(D_1, D_2)$ , and generate a codebook of size  $\lceil 2^{nR} \rceil$  at random by drawing each symbol of each codeword IID according to the marginal  $P_{\hat{X}}$ . To describe a source sequence  $x^n$ , the encoder tries to find a codeword  $\hat{x}^n(i)$  such that  $(x^n, \hat{x}^n(i)) \in \mathcal{T}_\epsilon^{(n)}(P_{X, \hat{X}})$  and then sends the smallest such index  $i$ . When observing  $i$ , the reconstructor produces  $\hat{x}^n(i)$ . Introduce a chance variable  $S$  that is one if the encoder was able to find at least one codeword  $\hat{x}^n(i)$  such that  $(x^n, \hat{x}^n(i)) \in \mathcal{T}_\epsilon^{(n)}(P_{X, \hat{X}})$ , and zero otherwise. The expected distortion for the random coding scheme can be bounded as

$$\begin{aligned} \mathbb{E}[d_1(X^n, \hat{X}^n)] &\stackrel{(i)}{=} \Pr[S = 0] \underbrace{\mathbb{E}[d_1(X^n, \hat{X}^n)|S = 0]}_{\leq d_{1,\max}} + \underbrace{\Pr[S = 1]}_{\leq 1} \underbrace{\mathbb{E}[d_1(X^n, \hat{X}^n)|S = 1]}_{\leq D_1(1+\epsilon)} \\ &\leq \Pr[S = 0] \cdot d_{1,\max} + D_1(1 + \epsilon) \end{aligned} \quad (3)$$

where (i) follows from the law of total expectation; the first underbraced inequality holds because we assume bounded distortions; the second underbraced inequality is trivial; the third underbraced inequality holds because  $S = 1$  implies that there exists a codeword  $\hat{x}^n(i)$  such that  $(x^n, \hat{x}^n(i)) \in \mathcal{T}_\epsilon^{(n)}(P_{X, \hat{X}})$ ; and for that codeword, we have  $d_1(x^n, \hat{x}^n) \leq D_1(1 + \epsilon)$  as a consequence of strong typicality because  $P_{\hat{X}_i|X_i}^*$  satisfies  $\mathbb{E}[d_1(X, \hat{X})] \leq D_1$ . Similarly, we obtain

$$\mathbb{E}[d_2(X^n, \hat{X}^n)] \leq \Pr[S = 0] \cdot d_{2,\max} + D_2(1 + \epsilon). \quad (4)$$

From the lecture we know that for  $R > R^I(D_1, D_2) = I(P_X, P_{\hat{X}_i|X_i}^*)$  and  $\epsilon$  small enough,  $\Pr[S = 0]$  tends to zero as  $n$  tends to infinity. Therefore, for an appropriate choice of  $\epsilon$  and  $n$  large enough, we obtain  $\mathbb{E}[d_1(X^n, \hat{X}^n)] \leq D_1 + \delta$  and  $\mathbb{E}[d_2(X^n, \hat{X}^n)] \leq D_2 + \delta$  from (3) and (4), respectively. Finally, we argue that we can get the same performance with a deterministic codebook. Note that the RHS of (3) and (4) depend only on  $\Pr[S = 0]$ . By the law of total probability, we have

$$\Pr[S = 0] = \sum_{\mathcal{C}} P(\mathcal{C}) \Pr[S = 0|\mathcal{C}],$$

so there must exist a deterministic codebook  $\mathcal{C}^*$  with  $\Pr[S = 0|\mathcal{C}^*] \leq \Pr[S = 0]$ , and this codebook also achieves  $\mathbb{E}[d_1(X^n, \hat{X}^n)] \leq D_1 + \delta$  and  $\mathbb{E}[d_2(X^n, \hat{X}^n)] \leq D_2 + \delta$ . (Since the codebook is generated independently of the source sequence,  $X^n$  is also distributed IID when conditioned on  $\mathcal{C}^*$ .)

- b) This result follows from the first part by defining  $\hat{\mathcal{X}} = \hat{\mathcal{X}}_1 \times \hat{\mathcal{X}}_2$ ,  $\hat{X} = (\hat{X}_1, \hat{X}_2)$  and setting  $\tilde{d}_1(x, \hat{x}) = d_1(x, \hat{x}_1)$  and  $\tilde{d}_2(x, \hat{x}) = d_2(x, \hat{x}_2)$ . The two reconstructors  $g_1^{(n)}$  and  $g_2^{(n)}$  can be viewed as the coordinates of a single reconstructor  $g^{(n)}: \{1, \dots, 2^{nR}\} \rightarrow \hat{\mathcal{X}}_1 \times \hat{\mathcal{X}}_2$ . Then,

$$\begin{aligned} \mathbb{E}[d_1(X^n, \hat{X}_1^n)] &\leq D_1 + \delta, \\ \mathbb{E}[d_2(X^n, \hat{X}_2^n)] &\leq D_2 + \delta \end{aligned}$$

if and only if

$$\begin{aligned} \mathbb{E}[\tilde{d}_1(X^n, \hat{X}^n)] &\leq D_1 + \delta, \\ \mathbb{E}[\tilde{d}_2(X^n, \hat{X}^n)] &\leq D_2 + \delta. \end{aligned}$$

Using Part a), we see that rates  $R > R(D_1, D_2)$  are achievable and rates  $R < R(D_1, D_2)$  are not achievable, where

$$R(D_1, D_2) = \min_{\substack{P_{\hat{X}|X}: \\ \mathbb{E}[\tilde{d}_1(X, \hat{X})] \leq D_1, \\ \mathbb{E}[\tilde{d}_2(X, \hat{X})] \leq D_2}} I(X; \hat{X}) = \min_{\substack{P_{\hat{X}_1, \hat{X}_2|X}: \\ \mathbb{E}[d_1(X, \hat{X}_1)] \leq D_1, \\ \mathbb{E}[d_2(X, \hat{X}_2)] \leq D_2}} I(X; \hat{X}_1, \hat{X}_2) = R^I(D_1, D_2).$$

- c) Let  $P_{\hat{X}|X}^*$  be a conditional PMF that achieves the minimum in the definition of  $R^I(D_1, D_2)$ . Let the conditional PMF  $\tilde{P}_{\hat{X}_1, \hat{X}_2|X}$  be given by

$$\tilde{P}_{\hat{X}_1, \hat{X}_2|X}(\hat{x}_1, \hat{x}_2|x) = P_{\hat{X}|X}^*(\hat{x}_1|x) \mathbb{I}\{\hat{x}_2 = \hat{x}_1\}.$$

Observe that  $\tilde{P}_{\hat{X}_1, \hat{X}_2|X}$  satisfies

$$\begin{aligned} \mathbb{E}[d_1(X, \hat{X}_1)] &\leq D_1, \\ \mathbb{E}[d_2(X, \hat{X}_2)] &\leq D_2. \end{aligned}$$

Consequently,

$$\begin{aligned} R^I(D_1, D_2) &= \min_{\substack{P_{\hat{X}_1, \hat{X}_2|X}: \\ \mathbb{E}[d_1(X, \hat{X}_1)] \leq D_1, \\ \mathbb{E}[d_2(X, \hat{X}_2)] \leq D_2}} I(P_X, P_{\hat{X}_1, \hat{X}_2|X}) \\ &\leq I(P_X, \tilde{P}_{\hat{X}_1, \hat{X}_2|X}) \\ &= I(P_X, P_{\hat{X}|X}^*) \\ &= R^I(D_1, D_2). \end{aligned}$$

**Problem 2**

**Source-Channel Separation with Feedback**

The first half of the proof proceeds along the same lines as the proof for the converse of the rate-distortion theorem; the second half of the proof is similar to the proof that feedback does not increase the capacity of a DMC. The claim is true because

$$\begin{aligned}
 k\mathbf{R}(\mathbf{D}) &\stackrel{(a)}{\leq} k\mathbf{R}(\mathbb{E}[d(U^k, \hat{U}^k)]) \\
 &\stackrel{(b)}{=} k\mathbf{R}\left(\sum_{i=1}^k \frac{1}{k} \mathbb{E}[d(U_i, \hat{U}_i)]\right) \\
 &\stackrel{(c)}{\leq} \sum_{i=1}^k \mathbf{R}(\mathbb{E}[d(U_i, \hat{U}_i)]) \\
 &\stackrel{(d)}{\leq} \sum_{i=1}^k I(U_i; \hat{U}_i) \\
 &= \sum_{i=1}^k [H(U_i) - H(U_i|\hat{U}_i)] \\
 &\stackrel{(e)}{=} H(U^k) - \sum_{i=1}^k H(U_i|\hat{U}_i) \\
 &\stackrel{(f)}{\leq} H(U^k) - \sum_{i=1}^k H(U_i|\hat{U}^k, U^{i-1}) \\
 &\stackrel{(g)}{=} H(U^k) - H(U^k|\hat{U}^k) \\
 &= I(U^k; \hat{U}^k) \\
 &\stackrel{(h)}{\leq} I(U^k; Y^n) \\
 &= H(Y^n) - H(Y^n|U^k) \\
 &\stackrel{(i)}{=} \sum_{i=1}^n [H(Y_i|Y^{i-1}) - H(Y_i|U^k, Y^{i-1})] \\
 &\stackrel{(j)}{\leq} \sum_{i=1}^n [H(Y_i) - H(Y_i|U^k, Y^{i-1})] \\
 &\stackrel{(k)}{=} \sum_{i=1}^n [H(Y_i) - H(Y_i|U^k, Y^{i-1}, X_i)] \\
 &\stackrel{(l)}{=} \sum_{i=1}^n [H(Y_i) - H(Y_i|X_i)] \\
 &= \sum_{i=1}^n I(X_i; Y_i) \\
 &\stackrel{(m)}{\leq} n\mathbf{C},
 \end{aligned}$$

where (a) follows from the monotonicity of  $\mathbf{R}(\mathbf{D})$ ; (b) holds because we use a single-letter distortion measure; (c) follows from Jensen's inequality and the convexity of  $\mathbf{R}(\mathbf{D})$ ; (d) follows from the definition of  $\mathbf{R}(\mathbf{D})$ ; (e) holds because  $U_1, \dots, U_k$  are independent; (f) holds because conditioning does not increase entropy; (g) follows from the chain rule; (h) follows from the data-processing inequality; (i) follows from the chain rule; (j) holds because conditioning does not increase entropy; (k) holds because  $X_i$  is a function of  $U^k$  and  $Y^{i-1}$ ; (l) holds because  $(U^k, Y^{i-1}) \text{ --- } X_i \text{ --- } Y_i$  form a Markov chain since we have a DMC; and (m) follows from the definition of  $\mathbf{C}$ .

### Problem 3

### Average Bit-Error Probability

Using the distortion function

$$d(u, \hat{u}) \triangleq I\{u \neq \hat{u}\} = \begin{cases} 0 & \text{if } u = \hat{u}, \\ 1 & \text{if } u \neq \hat{u}, \end{cases}$$

we see that

$$\frac{1}{k} \sum_{i=1}^k \Pr[U_i \neq \hat{U}_i] = \frac{1}{k} \sum_{i=1}^k \mathbb{E}[d(U_i, \hat{U}_i)] = \mathbb{E}[d(U^k, \hat{U}^k)] \geq D_{\min},$$

where  $D_{\min}$  denotes the smallest possible expected distortion. From Problem 2 we know that  $D_{\min}$  has to satisfy

$$R(D_{\min}) \leq \frac{n}{k} C.$$

Since we can always achieve an expected distortion of  $\frac{1}{2}$  (by setting  $\hat{U}_i$  to zero for example), we have  $D_{\min} \in [0, \frac{1}{2}]$ . Now,  $R(D_{\min}) = 1 - H_b(D_{\min})$  follows from the example given in the lecture (Bernoulli(1/2) source with Hamming distortion). Combining these results, we obtain

$$1 - H_b(D_{\min}) \leq \frac{n}{k} C.$$

By the assumption that  $\frac{k}{n} > C$  and by the monotonicity of the binary entropy function, we get

$$\frac{1}{k} \sum_{i=1}^k \Pr[U_i \neq \hat{U}_i] \geq D_{\min} \geq H_b^{-1}\left(1 - \frac{n}{k} C\right) > 0,$$

which shows that reliable communication above channel capacity is not possible (even with feedback, since the result from Problem 2 also holds if feedback is provided to the encoder).

(Without going into the details, note that this result is stronger than the converse for channel coding given in the lecture: There it was shown that the *block error probability*, i.e., the probability that at least one bit will not be decoded correctly, is bounded away from zero. Here you have shown that the *bit error rate* is bounded away from zero, i.e., that on average a constant fraction of the bits will not be decoded correctly.)