



Model Answers to Exercise 11 of May 11, 2017

<http://www.isi.ee.ethz.ch/teaching/courses/it2.html>

Problem 1

Method of Types

- a) To compute $|\mathcal{P}_n(\mathcal{X})|$, we make use of the following bijection between the set $\mathcal{P}_n(\mathcal{X})$ and the set of all sequences consisting of exactly n stars and $|\mathcal{X}| - 1$ separators: the number of stars at the beginning of a sequence corresponds to the number of occurrences of the first symbol from \mathcal{X} , the number of stars between the first and the second separator corresponds to the number of occurrences of the second symbol from \mathcal{X} , and so on. For example, the type $P_{\mathbf{x}} = (\frac{1}{5}, \frac{2}{5}, \frac{2}{5}) \in \mathcal{P}_5(\mathcal{X})$ corresponds to the sequence “*|**|**”. The number of all sequences consisting of exactly n stars and $|\mathcal{X}| - 1$ separators, and therefore $|\mathcal{P}_n(\mathcal{X})|$, is

$$\binom{n + |\mathcal{X}| - 1}{|\mathcal{X}| - 1}$$

because the sequences have length $n + |\mathcal{X}| - 1$ and separators have to appear at $|\mathcal{X}| - 1$ positions.

Here, $|\mathcal{P}_5| = \binom{7}{2} = 21$. This is much less than the upper bound

$$|\mathcal{P}_5| \leq (n + 1)^{|\mathcal{X}|} = 6^3 = 216.$$

The improved version of the upper bound is already closer:

$$|\mathcal{P}_5| \leq (n + 1)^{|\mathcal{X}| - 1} = 6^2 = 36.$$

- b) The type of $\mathbf{x} = (b, c, c, a, b)$ is $P_{\mathbf{x}}(a) = \frac{1}{5}$, $P_{\mathbf{x}}(b) = \frac{2}{5}$, $P_{\mathbf{x}}(c) = \frac{2}{5}$, i.e.,

$$P_{\mathbf{x}} = \left(\frac{1}{5}, \frac{2}{5}, \frac{2}{5} \right).$$

- c) The set of sequences having one particular type is called *type class*. In our case the size of the type class of $P_{\mathbf{x}}$ is

$$|\mathcal{T}^5(P_{\mathbf{x}})| = \frac{5!}{1!2!2!} = 30.$$

In order to compare this to the bounds derived in class we need to compute

$$H(P_{\mathbf{x}}) = -\frac{1}{5} \log \frac{1}{5} - 2 \cdot \frac{2}{5} \log \frac{2}{5} = \log 5 - \frac{4}{5} \log 2.$$

Hence,

$$|\mathcal{T}^5(P_{\mathbf{x}})| \leq e^{nH(P_{\mathbf{x}})} = e^{5 \log 5 - 4 \log 2} = \frac{5^5}{2^4} = \frac{3125}{16} \approx 195;$$

and

$$|\mathcal{T}^5(P_{\mathbf{x}})| \geq \frac{1}{(n+1)^{|\mathcal{X}|}} e^{nH(P_{\mathbf{x}})} = \frac{1}{216} \cdot \frac{3125}{16} \approx 0.90.$$

The improved lower bound is

$$|\mathcal{T}^5(P_{\mathbf{x}})| \geq \frac{1}{(n+1)^{|\mathcal{X}|-1}} e^{nH(P_{\mathbf{x}})} = \frac{1}{36} \cdot \frac{3125}{16} \approx 5.4.$$

These bounds are poor for small n . The lower bound can be improved further by using the exact number of types, $|\mathcal{P}_5| = 21$:

$$|\mathcal{T}^5(P_{\mathbf{x}})| \geq \frac{1}{|\mathcal{P}_5|} e^{nH(P_{\mathbf{x}})} = \frac{1}{21} \cdot \frac{3125}{16} \approx 9.3.$$

d) We already computed $H(P_{\mathbf{x}}) = \log 5 - \frac{4}{5} \log 2$. Next we compute

$$D(P_{\mathbf{x}}\|Q) = \frac{1}{5} \log \frac{1/5}{2/3} + 2 \cdot \frac{2}{5} \log \frac{2/5}{1/6} = \frac{1}{5} \log \frac{2^7 \cdot 3^5}{5^5}.$$

Then

$$\begin{aligned} Q^5(b, c, c, a, b) &= e^{-n(H(P_{\mathbf{x}}) + D(P_{\mathbf{x}}\|Q))} = e^{-5 \log 5 + 4 \log 2 - \log 2^7 - \log 3^5 + \log 5^5} \\ &= 2^{-3} \cdot 3^{-5} = \frac{1}{1944} \approx 5.14 \cdot 10^{-4}. \end{aligned}$$

Obviously, this could also have been computed directly as

$$Q^5(b, c, c, a, b) = \left(\frac{2}{3}\right)^1 \left(\frac{1}{6}\right)^2 \left(\frac{1}{6}\right)^2 = \frac{1}{2^3 \cdot 3^5} = \frac{1}{1944}.$$

e) Note that

$$\Pr[P_{\mathbf{X}} = P_{(b,c,c,a,b)}] = Q^5(\mathcal{T}^5(P_{(b,c,c,a,b)})).$$

We know from class that for $\mathbf{x} = (b, c, c, a, b)$

$$Q^5(\mathcal{T}^5(P_{\mathbf{x}})) \leq e^{-nD(P_{\mathbf{x}}\|Q)} = \frac{5^5}{2^7 \cdot 3^5} = \frac{3125}{31104} \approx 0.100 = 10.0 \cdot 10^{-2},$$

and

$$Q^5(\mathcal{T}^5(P_{\mathbf{x}})) \geq \frac{1}{(n+1)^{|\mathcal{X}|}} \cdot e^{-nD(P_{\mathbf{x}}\|Q)} = \frac{1}{216} \cdot \frac{3125}{31104} \approx 4.65 \cdot 10^{-4}.$$

This value can be improved as follows:

$$Q^5(\mathcal{T}^5(P_{\mathbf{x}})) \geq \frac{1}{(n+1)^{|\mathcal{X}|-1}} \cdot e^{-nD(P_{\mathbf{x}}\|Q)} = \frac{1}{36} \cdot \frac{3125}{31104} \approx 0.279 \cdot 10^{-2}.$$

Since all sequences in a type class have the same probability, the exact value of $Q^5(\mathcal{T}^5(P_{\mathbf{x}}))$ can be computed as

$$Q^5(\mathcal{T}^5(P_{\mathbf{x}})) = |\mathcal{T}^5(P_{\mathbf{x}})| e^{-nH(P_{\mathbf{x}}) - nD(P_{\mathbf{x}}\|Q)} = 30 \cdot \frac{1}{1944} = \frac{5}{324} \approx 1.54 \cdot 10^{-2}.$$

- a) We first treat some special cases. Note that $D(P\|Q)$ is finite if and only if $\text{supp}(P) \subseteq \text{supp}(Q)$.
- i) If $R_n \leq H(Q)$, then $\min_{P: H(P) \geq R_n} D(P\|Q) = D(Q\|Q) = 0$, so $P^* = Q$.
 - ii) If $R_n = \log \text{supp}(Q)$, then P^* is the uniform distribution over the support of Q , since this is the only distribution satisfying $\text{supp}(P) \subseteq \text{supp}(Q)$ and $H(P) \geq R_n$.
 - iii) If $R_n > \log \text{supp}(Q)$, then no P exists such that $D(P\|Q)$ is finite (and indeed, inspecting the proof for the universal source coding scheme shows that in this case, the probability of error is zero).

From now on, we assume $H(Q) < R_n < \log \text{supp}(Q)$. We use Lagrange multipliers to minimize $D(P\|Q)$ over P subject to $\sum_x P(x) = 1$ and $H(P) = r$. Let

$$J(P) = \sum_x P(x) \ln \frac{P(x)}{Q(x)} + \lambda_0 \left(\sum_x P(x) - 1 \right) + \lambda_1 \left(- \sum_x P(x) \ln P(x) - r \right).$$

Computing the partial derivative with respect to $P(\hat{x})$, we get

$$\frac{\partial J(P)}{\partial P(\hat{x})} = \ln \frac{P(\hat{x})}{Q(\hat{x})} + 1 + \lambda_0 - \lambda_1 \ln P(\hat{x}) - \lambda_1.$$

Setting this partial derivative to zero for every \hat{x} , we obtain

$$P^*(x) = \left(Q(x) e^{\lambda_1 - \lambda_0 - 1} \right)^{\frac{1}{1 - \lambda_1}},$$

which is, after substituting $\alpha = \frac{1}{1 - \lambda_1}$, equivalent to

$$P^*(x) = \frac{Q(x)^\alpha}{\sum_{x'} Q(x')^\alpha}.$$

Having determined the “form” of P^* , it remains to show how to choose α and to justify that P^* indeed achieves the minimum. For the interested reader, we now show a way to solve these issues. Rewrite P^* as

$$P^*(x) = \frac{Q(x)^{\frac{1}{1+\rho}}}{\sum_{x'} Q(x')^{\frac{1}{1+\rho}}} \tag{1}$$

and let $\rho > 0$ be such that $H(P^*) = R_n$. (The existence of such a ρ follows from the intermediate value theorem because $H(Q) < R_n < \log \text{supp}(Q)$, because $H(P^*)|_{\rho=0} = H(Q)$, because $\lim_{\rho \rightarrow \infty} H(P^*) = \log \text{supp}(Q)$, and because $H(P^*)$ is continuous in ρ .) Then, for every PMF P satisfying $H(P) \geq R_n$,

$$\begin{aligned} D(P\|Q) &= \sum_x P(x) \log \frac{P(x)}{Q(x)} \\ &= \sum_x P(x) \log \frac{P(x)^{1+\rho} \cdot P^*(x)^{1+\rho}}{P^*(x)^{1+\rho} \cdot P(x)^\rho \cdot Q(x)} \\ &= (1 + \rho)D(P\|P^*) + \rho H(P) + \sum_x P(x) \log \frac{P^*(x)^{1+\rho}}{Q(x)} \\ &\stackrel{(i)}{=} (1 + \rho)D(P\|P^*) + \rho H(P) + \log \gamma \end{aligned}$$

$$\begin{aligned}
&= (1 + \rho)D(P\|P^*) + \rho H(P) + \sum_x P^*(x) \log \frac{P^*(x)^{1+\rho}}{Q(x)} \\
&= (1 + \rho)D(P\|P^*) + \rho H(P) + \sum_x P^*(x) \log \frac{P^*(x)}{P^*(x)^{-\rho} \cdot Q(x)} \\
&= (1 + \rho)D(P\|P^*) + \rho H(P) - \rho H(P^*) + D(P^*\|Q) \\
&= (1 + \rho)D(P\|P^*) + \rho(H(P) - H(P^*)) + D(P^*\|Q) \\
&\stackrel{(ii)}{\geq} D(P^*\|Q),
\end{aligned}$$

where (i) holds because $\gamma = \frac{P^*(x)^{1+\rho}}{Q(x)} = \left\{ \sum_{x'} Q(x')^{\frac{1}{1+\rho}} \right\}^{-(1+\rho)}$ does not depend on x ; and (ii) holds because $\rho > 0$ and $H(P) \geq \mathbb{R}_n = H(P^*)$. Furthermore, equality in (ii) holds if and only if $P = P^*$. We conclude that in the case $H(Q) < \mathbb{R}_n < \log \text{supp}(Q)$, P^* given by (1) is the unique minimizer of $D(P\|Q)$ subject to $H(P) \geq \mathbb{R}_n$.

- b) To achieve $P_e^{(n)} \rightarrow 0$, $H(Q) < \mathbb{R}$ is sufficient (and we know that even if the encoder and the decoder know Q , the probability of error tends to one if $H(Q) > \mathbb{R}$). If $\mathbb{R} \geq 1$, then $P_e^{(n)} = 0$ is achieved for any binary distribution Q . If $\mathbb{R} \in (0, 1)$, then there exists a unique $p^* \in (0, \frac{1}{2})$ such that $H_b(p^*) = \mathbb{R}$ because $H_b(p)$ is strictly increasing in p for $p \in [0, \frac{1}{2}]$. Hence, $P_e^{(n)} \rightarrow 0$ is achieved if $Q = (q, 1 - q)$ satisfies $q < p^*$ or $q > 1 - p^*$.

Problem 3

Large Deviations

Using Sanov's theorem we know that the exponent we are looking for is

$$\inf_{f: \mathbb{E}_f[X^2] \geq \alpha^2} D(f\|\mathcal{N}(0, \sigma^2)). \quad (2)$$

If $\alpha^2 \leq \sigma^2$, then (2) is zero, which is achieved by choosing $f = \mathcal{N}(0, \sigma^2)$. From now on, assume that $\alpha^2 > \sigma^2$. For a fixed f satisfying $\mathbb{E}_f[X^2] \geq \alpha^2$, define $\beta^2 \triangleq \mathbb{E}_f[X^2] \geq \alpha^2 \geq \sigma^2$ and observe that

$$\begin{aligned}
D(f\|\mathcal{N}(0, \sigma^2)) &= \int_{-\infty}^{\infty} f(x) \ln \frac{f(x)}{\frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{x^2}{2\sigma^2}}} dx \\
&= -h(f) + \int_{-\infty}^{\infty} f(x) \ln \left(\sqrt{2\pi\sigma^2} e^{\frac{x^2}{2\sigma^2}} \right) dx \\
&= -h(f) + \ln \sqrt{2\pi\sigma^2} + \frac{1}{2\sigma^2} \int_{-\infty}^{\infty} f(x) x^2 dx \\
&= -h(f) + \frac{1}{2} \left(\ln 2\pi\sigma^2 + \frac{\beta^2}{\sigma^2} \right) \\
&\stackrel{(i)}{\geq} -\frac{1}{2} \ln 2\pi e \beta^2 + \frac{1}{2} \left(\ln 2\pi\sigma^2 + \frac{\beta^2}{\sigma^2} \right) \\
&= \frac{1}{2} \left(\frac{\beta^2}{\sigma^2} - 1 - \ln \frac{\beta^2}{\sigma^2} \right) \\
&\stackrel{(ii)}{\geq} \frac{1}{2} \left(\frac{\alpha^2}{\sigma^2} - 1 - \ln \frac{\alpha^2}{\sigma^2} \right) \text{ nats}, \quad (3)
\end{aligned}$$

where (i) holds because Gaussians maximize differential entropy for a fixed second moment; and (ii) holds because the expression in parentheses is monotonically increasing in β^2 for $\beta^2 \geq \sigma^2$ (this can be verified by taking the derivative with respect to β^2). Since (i) and (ii) hold with equality for the choice $f = \mathcal{N}(0, \alpha^2)$, the minimum in (2) in the case $\alpha^2 > \sigma^2$ is given by the RHS of (3).